

#### **Introduction to statistics** Lausanne, January 2025

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**Multiple Regression** 



What happens if both, age and weight variables were included in the same model ?

#### One multiple regression with two variables

Call: lm(formula = Height ~ Age + Weight, data = class)Residuals: Min 10 Median 30 Max -3.6248 -1.3016 -0.0176 0.8324 4.1019 Coefficients: Estimate Std. Error t value Pr(>|t|) (Intercept) 132.1943 5.0823 26.011 1.61e-14 \*\*\* 1.2267 0.5302 2.314 0.03431 \* Age Weight 0.2761 0.0695 3.973 0.00109 \*\* Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 2.255 on 16 degrees of freedom Multiple R-squared: 0.828, Adjusted R-squared: 0.8065 F-statistic: 38.52 on 2 and 16 DF, p-value: 7.646e-07

## This model allows us to determine the respective contribution of each variable <u>separately</u>.

Coefficients: Estimate Std. Error t value Pr(>|t|) (Intercept) 132.1943 5.0823 26.011 1.61e-14 \*\*\* Age 1.2267 0.5302 2.314 0.03431 \* Weight 0.2761 0.0695 3.973 0.00109 \*\* ---Signif. codes: 0 `\*\*\*` 0.001 `\*\*` 0.01 `\*` 0.05 `.` 0.1 ` ` 1

This is similar to the simple regression case.

Each test is conducted assuming that the tested parameter is the last one entering the model:

« If *weight* is already in the model, is the coefficient for *age* significantly different from 0? »

#### Two single regressions vs one multiple regression

Coefficients	:				
	Estimate	Std. Error t	value Pr(	>ltl)	
(Intercept)	142.57014	2.67989	53.200 <	2e-16 ***	
Weight	0.39523	0.05231	7.555 7.8	9e-07 ***	
Coefficients	s:				
	Estimate	Std. Error	t value P	r(>ltl)	
(Intercept)	125.2239	6.5217	19.201 5	.82e-13 *	**
Age	2.7871	0.4869	5.724 2	.48e-05 *	**
Coefficient	s:				
	Estimate	Std. Error	t value	Pr(>ltl)	
(Intercept)	132.1943	5.0823	26.011	1.61e-14	***
Age	1.2267	0.5302	2.314	0.03431	*
Weight	0.2761	0.0695	3.973	0.00109	**

While both age and weight seem significant by themselves, age is much less significant when weight is already included (see also the  $R^2$ ).

It is likely that a lot of the information provided by the age is also provided by the weight, so that there may be little need to have both terms in the model. 61

#### Multiple and adjusted R-squared

Multiple R-squared: 0.828, Adjusted R-squared: 0.8065

## As before, R<sup>2</sup> is the proportion of the total variance in the response data that is explained by the model.

Adding a new variable in the model will always increase R<sup>2</sup>, up to 1 when there the number of degrees of freedom is 0 (number of parameters to estimate = number of observations).

#### Multiple and adjusted R-squared

Multiple R-squared: 0.828, Adjusted R-squared: 0.8065

The adjusted R-squared adjusts for the number of variables in the model, and does not necessarily increase when the number of variables increase; it can even be negative.

 $R^2 = 1 - \frac{SS_{residuals}}{SS_{total}}$ 

It is always equal or below R<sup>2</sup>.

Adjusted R<sup>2</sup> = 
$$1 - \frac{SS_{residuals}}{SS_{total}} (n-K)$$

#### Example

# y <- rnorm(10) x1 <- rnorm(10); x2 <- rnorm(10); ...; x9 <rnorm(10)</pre>

summary(lm(y ~ x1)); summary(lm(y ~ x1+x2));

1:	Multiple	R-squared:	0.1419,
2:	Multiple	R-squared:	0.5173,
3:	Multiple	R-squared:	0.557,
4:	Multiple	R-squared:	0.5577,
5:	Multiple	R-squared:	0.7953,
6:	Multiple	R-squared:	0.8321,
7:	Multiple	R-squared:	0.984,
8:	Multiple	R-squared:	0.9851,
g .	Multiple	R-squared.	1

Adjusted	R-squared:	0.03464
Adjusted	R-squared:	0.3794
Adjusted	R-squared:	0.3355
Adjusted	R-squared:	0.2039
Adjusted	R-squared:	0.5395
Adjusted	R-squared:	0.4962
Adjusted	R-squared:	0.9281
Adjusted	R-squared:	0.866
Adjusted	R-squared:	NaN

#### The last regression from the example

Call:					
lm(formula	= y ~ x1 +	x2 + x3 -	+ x4 + x5	+ x6 + x7	+ x8 + x9
Residuals:					
ALL 10 res	iduals are	0: no res:	ldual deg	rees of fre	eedom!
Coefficien	ts:				
	Estimate	Std. Erro	r t value	Pr(> t )	
(Intercept	) -0.02693	N	A NA	NA	
x1	0.53886	N	A NA	NA	
x2	-0.52227	N	A NA	NA	
хЗ	0.51881	N	A NA	NA	
x4	0.74757	N	A NA	NA	
x5	0.14394	N	A NA	NA	
хб	-0.65387	N	A NA	NA	
x7	-0.48271	N	A NA	NA	
x8	-0.62487	N	A NA	NA	
x9	0.23759	NZ	A NA	NA	
Residual s	tandard <u>err</u>	or: NaN <u>o</u> 1	n O deg <u>re</u> e	es of fr <u>ee</u> d	dom
Multiple R	-squared:	1,	Adjuste	d R-square	d: NaN

F-statistic: NaN on 9 and 0 DF, p-value: NA

#### F-statistic for significance of regression

Coefficients	5						
	Estimate	Std.	Error	t	value	Pr(> t )	
(Intercept)	81.77355	12.	90896		6.335	9.92e-06	* * *
Age	3.11575	1.	34668		2.314	0.03431	*
Weight	0.35064	0.	08827		3.973	0.00109	* *
F-statistic:	38.52 or	n 2 an	d 16 D	)F,	p-va	alue: 7.64	6e-07

Again, the F-statistic allows us to test if the whole regression (adding all variables *vs* having only the intercept in) is significant.

If any of the tests for the individual variables is significant, the Ftest will generally be significant as well.

However, even if no individual variable is significant (e.g. p < 0.05), the F-test can still be significant.

### Categorical variables, dummy variables and contrasts

#### Categorical variables

We'd like to use categorical variables in a linear model, as in:

#### Height = $b_0 + b_1$ Age + $b_2$ « Gender » + error

Intuitively, we want to estimate a « Male » and a « Female » effect.

#### Categorical variables

We'd like to use categorical variables in a linear model, as in:

#### $Height = b_0 + b_1 Age + b_2 \ll Gender \gg + error$

Intuitively, we want to estimate a « Male » and a « Female » effect.

In practice, categorical variables (factors in R) are turned (by default, based on alphabetical order) into **dummy variables** of the form

Gender = 
$$\begin{cases} 1 \text{ if Female} \\ 2 \text{ if Male} \end{cases}$$

#### Example of summary results of the lm command in R

Call: lm(formula = Height ~ Age + Gender, data = class)Residuals: 10 Median 30 Min Max -3.483 -1.910 -0.319 1.326 5.317 Coefficients: Estimate Std. Error t value Pr(>|t|) (Intercept) 124.5241 5.8886 21.147 4.04e-13 \*\*\* Age 2.7276 0.4398 6.202 1.27e-05 \*\*\* GenderM 2.8362 1.2797 2.216 0.0415 \* Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1 Residual standard error: 2.78 on 16 degrees of freedom Multiple R-squared: 0.7387, Adjusted R-squared: 0.706

F-statistic: 22.61 on 2 and 16 DF, p-value: 2.176e-05

Example of summary results of the lm command in R

Call: lm(formula = Height ~ Age + Gender, data = class)Residuals: Min 10 Median 30 Max baseline for -3.483 -1.910 -0.319 1.326 5.317 height among Female Coefficients: Estimate Std. Error t value Pr(>|t|) 5.8886 21.147 4.04e-13 \*\*\* (Intercept) 124.5241 2.7276 0.4398 6.202 1.27e-05 \*\*\* Age GenderM 2.8362 1.2797 2.216 0.0415 \* 0 (\*\*\*\* 0.001 (\*\*\* 0.01 (\*\* 0.05 (.' 0.1 (' 1 Signif. codes:

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The factor GenderM corresponds to the difference in baseline for Males compared to females

#### Graphical interpretation

The model specifies 2 straight lines, with the same slope but different yintercepts:



class\$Age

What if we don't use a linear model?

## We could also compute the difference in means between males and females directly:

# This result is slightly different from the 2.8362 cm difference found with the linear model.

Where does the difference come from ?

So far, we have assumed a difference between the lines, but the same slope; that is, for both men and women, the effect of age is the same.

If this assumption is incorrect, it means that there is an *interaction* between the factors « age » and « gender », that is, the effect of age is different depending on the gender.

Interactions are modeled in R in the following way:

Im(formula = Height ~ Age + Gender + Age:Gender)

which is equivalent to

Im(formula = Height ~ Age \* Gender)



#### Coefficients with an interaction

Call: lm(formula = Height ~ Age \* Gender, data = class)Residuals: 10 Median 30 Min Max -3.4429 -1.7844 -0.3648 1.3730 5.3571 Coefficients: Estimate Std. Error t value Pr(>|t|) (Intercept) 122.1500 9.6409 12.670 2.05e-09 \*\*\* 2.9071 0.7256 4.007 0.00114 \*\* Age GenderM 6.7443 12.4109 0.543 0.59483 Age:GenderM -0.2940 0.9285 -0.317 0.75585 \_ \_ \_ Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1 Residual standard error: 2.862 on 15 degrees of freedom Multiple R-squared: 0.7404, Adjusted R-squared: 0.6885 F-statistic: 14.26 on 3 and 15 DF, p-value: 0.0001152

#### The coefficients can be interpreted as follows:

#### According to the model, the *height* is equal to

122.15 (the intercept) plus 2.9071 times the person's age plus 6.7443, but only for males -0.2940 times the person's age, but only for males.

#### **Different** slopes



#### No interaction

With interaction

> model <- lm( Height ~ Age+Gender1, data=class)</pre> > summary(model) Call: lm(formula = Height ~ Age + Gender1, data = class)Residuals: Min 10 Median 30 Max -3.483 -1.910 -0.319 1.326 5.317 Coefficients: Estimate Std. Error t value Pr(>|t|) (Intercept) 127.3603 5.9587 21.374 3.43e-13 \*\*\* 2.7276 0.4398 6.202 1.27e-05 \*\*\* Age Gender1F -2.8362 1.2797 -2.216 0.0415 \* \_ \_ \_ Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1 Residual standard error: 2.78 on 16 degrees of freedom Multiple R-squared: 0.7387, Adjusted R-squared: 0.706 F-statistic: 22.61 on 2 and 16 DF, p-value: 2.176e-05 > model <- lm( Height ~ Age+Gender, data=class)</pre> > summary(model) Call: lm(formula = Height ~ Age + Gender, data = class)Residuals: Min 1Q Median 30 Max -3.483 -1.910 -0.319 1.326 5.317 Coefficients: Estimate Std. Error t value Pr(>|t|) (Intercept) 124.5241 5.8886 21.147 4.04e-13 \*\*\* 2.7276 0.4398 6.202 1.27e-05 \*\*\* Age GenderM 2.8362 1.2797 2.216 0.0415 \* \_\_\_ Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1 Residual standard error: 2.78 on 16 degrees of freedom Multiple R-squared: 0.7387, Adjusted R-squared: 0.706

F-statistic: 22.61 on 2 and 16 DF, p-value: 2.176e-05

The two models are exactly the same; only the way we look at the coefficient changes.

Gender1 <- relevel(Gender, ref="M")</pre>

#### What if Males were the baseline ?

### **Diagnostic tools**

It is always possible to fit a linear model and find a slope and intercept ... but it does not mean that the model is meaningful !

Examination of *residuals*: (which should show no obvious trend, since any systematic effect in the residuals should ideally be captured by the model):

- Normality
- Time effects
- Nonconstant variance Curvature

#### **Examination of** *residuals*



Works only for simple regression (only one variable on x axis)

Works also for multiple regression

*High leverage* ('influential') points are far from the center, and have potentially greater influence

One way to assess points is through the *hat values* (obtained from the *hat matrix H*):

$$\hat{y} = Xb = X(X'X)^{-1}X'y = Hy$$
  
 $h_i = \Sigma_j h_{ij_2}$ 

Average value of h = number of coefficients/n (including the intercept) = p/n

Cutoff typically 2p/n or 3p/n



Hat values

Actual fit

```
>hat <- lm.influence( model )
>plot( hat$hat )
>abline(h=c(c(2,3)*2/19),lty=c(2,3),col=c("blue","red") )
```



Narrow bands:describe the uncertainly about the regression lineWide bands:describe where most (95% by default) predictions would fall,<br/>assuming normality and constant variance.

In R: ?predict.lm